Statistical inference for rough volatility: Minimax theory

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• This presentation is based on a joint work with Carsten Chong, Marc Hoffmann, Yanghui Liu, Mathieu Rosenbaum :



- Statistical inference for rough volatility : Minimax Theory, arXiv, 2022.
- Statistical inference for rough volatility : Central limit theorems, arXiv, 2022.

Statistical model

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Rough volatility model

We consider an asset whose log price is modelled by a stochastic process ${\cal S}$ with dynamic

$$\begin{cases} dS_t = \sigma_t dB_t, \\ \sigma_t^2 = \exp(\eta W_t^H) \end{cases}$$

where $(H, \eta) \in \mathcal{D} = [H_-, H_+] \times [\eta_-, \eta_+]$, and where *B* is a Brownian motion and W^H is a Fractional Brownian Motion independent of *B*.

High frequency asymptotic

We aim at estimating the parameters H and η from high frequency observations of S. More precisely, we observe S at times $\frac{i}{n}$ with i = 0, ..., n and we suppose that $n = 2^N$.

Definition

A fractional Brownian motion W^H with Hurst index H with 0 < H < 1 is the Gaussian process satisfying

$$Cov(W_t^H, W_s^H) = \frac{1}{2}(t^{2H} + s^{2H} - |t - s|^{2H}).$$

Proposition

- W^H has stationary increments.
- $(a^{-H}W_{at}^{H})_{t}$ is a fractional Brownian motion with Hurst index H.

- Estimation in stationary Gaussian sequences : Fox & Taqqu '86 and '87; Istas & Lang '97; Coeurjolly '01.
- Estimation of the Hurst parameter in additive noise models : Gloter and Hoffmann '07 when H>1/2 only.
- Estimation of the Hurst parameter in multiplicative noise models : Rosenbaum '08 when H>1/2 only.
- LAN property : Kawai '13; Brouste, Fukasawa '18.
- Whittle estimators : Fukasawa, Takabatake, Westphal '19

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Definition

The rate v_n is said to be a **lower rate of convergence** over \mathcal{D} for estimating H if there exists c > 0 such that

$$\liminf_{n\to\infty}\inf_{\widehat{H}_n}\sup_{(H,\eta)\in\mathcal{D}}\mathbb{P}^n_{H,\eta}(v_n^{-1}|\widehat{H}_n-H|\geq c)>0$$

where the infimum is taken over all admissible estimators.

Theorem

The rate $v_n(H) = n^{-1/(4H+2)}$ is a lower rate of convergence for estimating H over \mathcal{D} .

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Construction of the estimator

• Our observations are given by

$$S_{i/n}=\int_0^{i/n}\sigma_s\,dB_s.$$

Stochastic calculus gives

$$(S_{i/n} - S_{(i-1)/n})^2 = \int_{(i-1)/n}^{i/n} \sigma_t^2 dt \ X_{i,n}^2$$

where $X_{i,n}$ are independent standard Gaussian variables.

• Using the logarithm gives

$$\log\left(n\left(S_{i/n}-S_{(i-1)/n}\right)^{2}\right) = \log\left(n\int_{(i-1)/n}^{i/n}\sigma_{t}^{2}\,dt\right) + \log\left(X_{i,n}^{2}\right).$$

• Here, we want to use the approximation

$$n\int_{(i-1)/n}^{i/n}\sigma_t^2 dt \approx \sigma_{i/n}^2 = \exp(\eta W_{i/n}^H)$$

so that we would be down to

$$Y_{i,n} = \eta W_{i/n}^H + \varepsilon_{i,n}$$

where the additive noise $\varepsilon_{i,n}$ is a well-behaved noise process.

Definition (Energy Level)

For $j \leq N$, we define the energy levels of the Fractional Brownian trajectory by

$$Q_j = \eta^2 2^{-j} \sum_{k=0}^{2^{j-1}} (W_{(k+1)2^{-j}}^H - W_{k2^{-j}}^H)^2.$$

Proposition

$$\mathbb{E}[(Q_j - rac{1}{2}\eta^2 2^{-2jH})^2] \leq C 2^{-j(1+4H)}.$$

Energy levels

• Therefore

$$2^{-2H} pprox rac{Q_{j+1}}{Q_j}$$

and we shall estimate H through this ratio.

• Therefore, we need to estimate correctly the energy levels, and we start by estimating the increments

$$d_{j,k} = \eta 2^{-j/2} (W^{H}_{(k+1)2^{-j}} - W^{H}_{k2^{-j}}).$$

• This is done by averaging the observations of the fractional Brownian motion to dampen the effects of the noise, which gives

$$\widetilde{d}_{j,k} = 2^{-j/2} \sum_{l=0}^{2^{N-j}-1} (Y_{(k+1)2^{N-j}+l,n} - Y_{k2^{N-j}+l}).$$

• *d* can be decomposed as

$$\begin{split} \widetilde{d}_{j,k} &= 2^{-j/2} \sum_{l=0}^{2^{N-j}-1} \left(Y_{(k+1)2^{N-j}+l,n} - Y_{k2^{N-j}+l} \right) \\ &= \underbrace{2^{-j/2} \eta \sum_{l=0}^{2^{N-j}-1} \left(W_{(k+1)2^{-j}+(l+1)/n}^{H} - W_{(k+1)2^{-j}+l/n}^{H} \right)}_{\approx d_{j,k}} \\ &+ \underbrace{2^{-j/2} \sum_{l=0}^{2^{N-j}-1} \left(\varepsilon_{(k+1)2^{N-j}+l,n} - \varepsilon_{k2^{N-j}+l} \right)}_{=:e_{j,k}}. \end{split}$$

- Provided ε is well behaved, $d_{j,k}$ is a good estimator for $d_{j,k,n}$.
- We need to correct the effect of the second moment of the noise in

$$\widetilde{d}_{j,k}^2 pprox (d_{j,k,n})^2 + 2d_{j,k,n}e_{j,k} + e_{j,k}^2,$$

and therefore we write

$$\widehat{Q}_{j,n} = \sum_{k=0}^{2^{j-1}} \widetilde{d}_{j,k}^2 - \mathbb{E}[e_{j,k}^2].$$

• Again, provided ε is well behaved, we have

$$\mathbb{E}[(\widehat{Q}_{j,n}-Q_j)^2] \leq C2^j/n^2.$$

• Since $2^{-2H} \approx Q_{j+1}/Q_j$, our final estimator of H is

$$-rac{1}{2}\log\Big(rac{\widehat{Q}_{j+1,n}}{\widehat{Q}_{j,n}}\Big)$$

but we still need to find an appropriate level j.

- We do a Biais-variance decomposition, the biais is of order $2^{-j/2}$ while the variance is of order $n^{-1}2^{j(4H+1)/2}$.
- Balancing both term yields $2^j \approx n^{1/(2H+1)}$

Adaptive estimation

We can cook up $\widehat{J}_n^* \approx n^{1/(2H+1)}$ by choosing \widehat{J}_n^* so that

$$Q_{J_n^*,n} \approx n^{-1} 2^{J_n^*}$$

The estimator of H is eventually given by

$$\widehat{H}_n = -rac{1}{2}\lograc{\widehat{Q}_{\widehat{J}_n^*+1}}{\widehat{Q}_{\widehat{J}_n^*}}.$$

Theorem

 $n^{1/(4H+2)}(\widehat{H}_n - H)$ is bounded in probability uniformly over \mathcal{D} .

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Unfortunately, when H < 1/2, The procedure detailed previously does not work because ε is not nice...

$$\begin{split} \eta W_{i/n}^{H} + \varepsilon_{i,n} &= \log \left(n \int_{(i-1)/n}^{i/n} \sigma_t^2 \, dt \right) + \log \left(X_{i,n}^2 \right) \\ &\approx \log \left(\sigma_{(i-1)/n}^2 \right) + \log \left(X_{i,n}^2 \right). \end{split}$$

Proposition

There exists random variable Z_0 bounded in $L^2(\mathbb{P}_{H,\eta})$ uniformly on \mathcal{D} such that

$$\log\left(n\int_{i/n}^{(i+1)/n} \sigma_u^2 du\right) = \cdots$$

$$= \sum_{b=2}^{2S} \sum_{s=1}^{2S} \frac{(-1)^{s-1}}{s} \sum_{\substack{\mathbf{r} \in \{1,\dots,S\}^s \\ \sum_j \mathbf{r}_j = b}} \prod_{j=1}^s \frac{\eta^{\mathbf{r}_j}}{\mathbf{r}_j!} \frac{1}{n} \int_{i/n}^{(i+1)/n} (W_u^H - W_{i/n}^H)^{\mathbf{r}_j} du$$

$$+ n \int_{i/n}^{(i+1)/n} \eta W_u^H du + Z(i,n) \cdot n^{-H^*(S+1)}$$

where the random variables Z(i, n) satisfy $|Z(i, n)| \leq Z_0$.

- We can still define the energy levels associated with these observations, but the scaling is hugely influenced by the presence of the additional terms.
- Indeed, the we cannot hope having an expression like $\mathbb{E}[(Q_{j,n} \eta^2 \kappa(H) 2^{-2jH})^2] \leq C 2^{-j(1+4H)} \text{ because the terms } (W_u^H W_{i/n}^H)^{\mathbf{r}_j} \text{ create additional scaling terms of order } 2^{2Hj}.$

Proposition

There exist explicit functions of H denoted κ_a such that if $S \ge 1/(4H_-) + 1/2$ and $S > H_+/(2H_-) - 1/2$, we have

$$\mathbb{E}_{H,\eta}\Big[\big(Q_j-\sum_{a=1}^{S}\eta^{2a}2^{-2aHj}\kappa_a(H)\big)^2\Big]\leq C2^{-j(1+4H)}$$

for some constant C depending only on S.

Energy levels

- Therefore the scaling $Q_{j+1}/Q_j = 2^{-2H}$ is no longer exact.
- Instead we have

$$\frac{Q_{j+1}}{Q_j} \approx \frac{\sum_{a=1}^{S} \eta^{2a} 2^{-2aH(j+1)} \kappa_a(H)}{\sum_{a=1}^{S} \eta^{2a} 2^{-2aHj} \kappa_a(H)} \approx 2^{-2H} + O(2^{-2Hj})$$

• We can build an estimator using the same procedure as in the simplified model, but it will exhibit a slower rate of convergence because of the additional contribution of order 2^{-2Hj} in the biais of this estimator.

Theorem

 $(n^{1/(4H+2)} \wedge n^{2H})(\hat{H}_n - H)$ is bounded in probability uniformly over \mathcal{D} .

- We need an additional biais correction procedure to improve the convergence rate of this estimator.
- First we need an estimator for η .

Proposition

If $v_n(\widehat{H}_n - H)$ is bounded in probability uniformly over \mathcal{D} , we can build an estimator $\widehat{\eta}_n$ of η such that $v_n \log(n)^{-1}(\widehat{\eta}_n - \eta)$ is bounded in probability uniformly over \mathcal{D} .

• The scaling of the energy levels is given by $Q_j \approx \sum_{a=1}^{S} \eta^{2a} 2^{-2aHj} \kappa_a(H)$ and we want to cut this sum to a = 1.

• Thus we need to replace Q_j by

$$Q_j - \sum_{a=2}^{S} \eta^{2a} 2^{-2aHj} \kappa_a(H) \approx \eta^2 2^{-2Hj} \kappa_1(H).$$

• So we replace \widehat{Q}_j by $\widehat{Q}_j^c(\widehat{H}_n, \widehat{\eta}_n)$

$$\widehat{Q}_{j}^{c}(\widetilde{H},\widetilde{\eta}) = \widehat{Q}_{j} - \sum_{a=2}^{S} \widetilde{\eta}^{2a} 2^{-2a\widetilde{H}j} \kappa_{a}(\widetilde{H})$$

Notice that

$$\frac{\widehat{Q}_j^c(H,\eta)}{\widehat{Q}_j^c(H,\eta)} = 2^{-2H}$$

as in the simplified setup.

• Thus we define

$$\widehat{H}_n^{(1)} = -rac{1}{2}\lograc{\widehat{Q}_{J_n^*+1}(\widehat{H}_n,\widehat{\eta}_n)}{\widehat{Q}_{J_n^*}(\widehat{H}_n,\widehat{\eta}_n)}$$

• However, we cannot immediately retrieve the convergence rate of the simplified setup because $\widehat{Q}_{j}^{c}(\widehat{H}_{n},\widehat{\eta}_{n}) \neq \widehat{Q}_{j}^{c}(H,\eta)$. The difference $\widehat{Q}_{j}^{c}(\widehat{H}_{n},\widehat{\eta}_{n}) - \widehat{Q}_{j}^{c}(H,\eta)$ also creates a biais that we can still control.

Theorem

 $(n^{1/(4H+2)} \wedge n^{4H(H+1)/(2H+1)})(\widehat{H}_n^{(1)} - H)$ is bounded in probability uniformly over \mathcal{D} .

• We repeat the biais correction procedure by defining a sequence of estimators

$$\widehat{H}_n^{(m+1)} = -\frac{1}{2} \log \frac{\widehat{Q}_{J_n^*+1}(\widehat{H}_n^{(m)}, \widehat{\eta}_n^{(m)})}{\widehat{Q}_{J_n^*}(\widehat{H}_n^{(m)}, \widehat{\eta}_n^{(m)})}$$

Theorem

 $(n^{1/(4H+2)} \wedge n^{2H(2H+m+1)/(2H+1)})(\widehat{H}_n^{(m)} - H)$ is bounded in probability uniformly over \mathcal{D} .

• We take $m_{opt} > m > 1/(4H) - 2H - 1$ for any $H_- < H < H_+$ and we get the same convergence rate $n^{1/(4H+2)}$ as in the simplified model.

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- The lower bound for estimating the parameter H is $n^{-1/(4H+2)}$, and we have developed an estimator achieving this rate.
- This rate is unusual and could seem counter-intuitive at first glance since it the optimal rate for estimating β -Holder continuous function in most models is usually $n^{-\beta/(2\beta+1)}$.
- But we do not seek to reconstruct the spot volatility.
- The rougher the volatility is, the easier it is to see it.
- Beyond this simplified framework, we can extend these idea to (rough) stochastic Volterra differential equations.

Thank you!

Asymptotic behaviour

• These Estimators are asymptotically Gaussian...



• In practice, the models considered in finance do not use Fractional Brownian motion but non-parametric approximations of the fractional Brownian motion :

$$X_t = X_0 + \int_0^t b_s \, ds + \int_0^t \sigma_s \, dB_s$$

$$\sigma_t^2 = \sigma_0^2 + \int_0^t a_s \, ds + \int_0^t g(t-s)\eta_s \, d\widetilde{B}_s$$

where $g(t) \approx t^{H-1/2}$.

• We can extend our estimators to embrace this framework.