

A GMM Approach to Estimate the Roughness of Stochastic Volatility

Mikko Pakkanen

Department of Mathematics Imperial College London

Minisymposium: Efficient Inference for Large and High-Frequency Data ICIAM 2023 Tokyo, 25 August 2023

Joint work with Anine Bolko, Kim Christensen and Bezirgen Veliyev (Aarhus)

Imperial College London



Agenda

• We develop an estimation/inference method for log-normal spot volatility models based on realised measures.



- We develop an estimation/inference method for log-normal spot volatility models based on realised measures.
- We employ the generalised method of moments (GMM).



- We develop an estimation/inference method for log-normal spot volatility models based on realised measures.
- We employ the generalised method of moments (GMM).
- Whilst the approach is fully generic, we focus on models that lend themselves to rough volatility.



- We develop an estimation/inference method for log-normal spot volatility models based on realised measures.
- We employ the generalised method of moments (GMM).
- Whilst the approach is fully generic, we focus on models that lend themselves to rough volatility.
- We take seriously the intrinsic sources of bias integration effect and measurement error.



- We develop an estimation/inference method for log-normal spot volatility models based on realised measures.
- We employ the generalised method of moments (GMM).
- Whilst the approach is fully generic, we focus on models that lend themselves to rough volatility.
- We take seriously the intrinsic sources of bias integration effect and measurement error.
- We derive asymptotic theory for the GMM estimator and assess its finite-sample properties in a simulation experiment.



- We develop an estimation/inference method for log-normal spot volatility models based on realised measures.
- We employ the generalised method of moments (GMM).
- Whilst the approach is fully generic, we focus on models that lend themselves to rough volatility.
- We take seriously the intrinsic sources of bias integration effect and measurement error.
- We derive asymptotic theory for the GMM estimator and assess its finite-sample properties in a simulation experiment.
- Applying the method to equity index data, we study whether spot volatility is best described by a rough process.

Measuring the roughness of volatility

Generalised method of moments approach

Epilogue



Scaling of fractional Brownian motion

The *q*-th order variogram of fractional Brownian motion (fBm), for q > 0, scales as follows:

$$u(q,\Delta) \mathrel{\mathop:}= \mathbb{E}[|W^H_{t+\Delta} - W^H_t|^q] \propto \Delta^{qH}, \quad \Delta > 0$$

Scaling of fractional Brownian motion

The *q*-th order variogram of fractional Brownian motion (fBm), for q > 0, scales as follows:

$$\mathbf{v}(q,\Delta) \mathrel{\mathop:}= \mathbb{E}[|W^H_{t+\Delta} - W^H_t|^q] \propto \Delta^{qH}, \quad \Delta > 0$$

Thus,

$$\frac{\log v(q,\Delta)}{q} = \operatorname{const} + H \log \Delta.$$



Estimation of H

For daily realised variance RV_k on days k = 0, 1, ..., N, compute the empirical variogram

$$m(q,\Delta) = \frac{1}{N} \sum_{k=1}^{N} |\log(\mathbb{R}V_{k\Delta}) - \log(\mathbb{R}V_{(k-1)\Delta})|^{q}, \quad \Delta = 1, \ldots, \lfloor \frac{N}{\Delta} \rfloor.$$



Estimation of H

For daily realised variance RV_k on days k = 0, 1, ..., N, compute the empirical variogram

$$m(q,\Delta) = \frac{1}{N} \sum_{k=1}^{N} |\log(\mathbb{R}V_{k\Delta}) - \log(\mathbb{R}V_{(k-1)\Delta})|^{q}, \quad \Delta = 1, \ldots, \lfloor \frac{N}{\Delta} \rfloor.$$

To estimate H, Gatheral, Jaisson and Rosenbaum (2018) suggest we...

regress
$$\frac{\log m(q,\Delta)}{q}$$
 on $\log \Delta$.

GMM approach



Roughness of Nikkei 225 volatility





General semimartingale model

$$\frac{\mathrm{d}S_t}{S_t} = \mu_t \mathrm{d}t + \sigma_t \mathrm{d}W_t + J_t, \quad t \ge 0$$



General semimartingale model

$$\frac{\mathrm{d}S_t}{S_t} = \mu_t \mathrm{d}t + \sigma_t \mathrm{d}W_t + J_t, \quad t \ge 0$$

Drift part

• Drift process $(\mu_t)_{t\geq 0}$



General semimartingale model

$$\frac{\mathrm{d}S_t}{S_t} = \mu_t \mathrm{d}t + \sigma_t \mathrm{d}W_t + J_t, \quad t \ge 0$$

Drift part

• Drift process $(\mu_t)_{t\geq 0}$

Diffusive part

- Spot volatility process (σ_t)_{t≥0}
- Standard Brownian motion $(W_t)_{t\geq 0}$



General semimartingale model

$$\frac{\mathrm{d}S_t}{S_t} = \mu_t \mathrm{d}t + \sigma_t \mathrm{d}W_t + J_t, \quad t \ge 0$$

Drift part

• Drift process $(\mu_t)_{t\geq 0}$

Diffusive part

- Spot volatility process $(\sigma_t)_{t\geq 0}$
- Standard Brownian motion $(W_t)_{t\geq 0}$

Jump part [omitted in what follows]

• Jump process (e.g., Lévy process) $(J_t)_{t\geq 0}$



Roughness measured via proxy?

Is it okay to think that $RV_k^n \approx \sigma_k^2$ to infer that σ is rough too?



Roughness measured via proxy?

Is it okay to think that $RV_k^n \approx \sigma_k^2$ to infer that σ is rough too?

Bias 1: Integration

$$\sigma^2 \longrightarrow IV_k := \int_{k-1}^k \sigma_t^2 dt$$
 biases \widehat{H} upward



Roughness measured via proxy?

Is it okay to think that $RV_k^n \approx \sigma_k^2$ to infer that σ is rough too?

Bias 1: Integration

$$\sigma^2 \longrightarrow IV_k := \int_{k-1}^k \sigma_t^2 \mathsf{d} t$$
 biases \widehat{H} upward

• Integration is a smoothing operation.



Roughness measured via proxy?

Is it okay to think that $RV_k^n \approx \sigma_k^2$ to infer that σ is rough too?

Bias 1: Integration

$$\sigma^2 \longrightarrow IV_k \coloneqq \int_{k-1}^k \sigma_t^2 \mathsf{d}t \qquad \mathsf{biases} \ \widehat{H} \ \mathsf{upward}$$

- Integration is a smoothing operation.
- Gatheral, Jaisson and Rosenbaum (2018) compute an estimate of the bias.

GMM approach



Is spot volatility rough?

Bias 2: Measurement error

$$IV_k \longrightarrow RV_k^n = IV_k + \underbrace{e_k^n}_{\text{error}}$$

biases \widehat{H} downward



Bias 2: Measurement error

$$VV_k \longrightarrow RV_k^n = IV_k + \underbrace{e_k^n}_{\text{error}}$$

biases \widehat{H} downward

 Asymptotic theory suggests that eⁿ_k, k = 1, 2, ..., are approximately uncorrelated, so they appear as noise.



Bias 2: Measurement error

$$V_k \longrightarrow RV_k^n = IV_k + \underbrace{e_k^n}_{\text{error}}$$

biases \widehat{H} downward

- Asymptotic theory suggests that eⁿ_k, k = 1, 2, ..., are approximately uncorrelated, so they appear as noise.
- Noise looks like roughness, "illusive scaling" (Fukasawa, Takabatake and Westphal, 2022).



Bias 2: Measurement error

$$V_k \longrightarrow RV_k^n = IV_k + \underbrace{e_k^n}_{\text{error}}$$

biases \widehat{H} downward

- Asymptotic theory suggests that eⁿ_k, k = 1, 2, ..., are approximately uncorrelated, so they appear as noise.
- Noise looks like roughness, "illusive scaling" (Fukasawa, Takabatake and Westphal, 2022).
- Mitigation via non-linear least squares (Bennedsen, Lunde and Pakkanen, 2022) or quasi-likelihood estimation (Fukasawa, Takabatake and Westphal, 2022).

Measuring the roughness of volatility

Generalised method of moments approach

Epilogue

Measuring roughness

GMM approach

Epilogue 000

Probabilistic setup

Log-normal spot volatility model

$$\sigma_t^2 := \xi \exp\left(\frac{\mathbf{Y}_t - \frac{1}{2}\kappa(\mathbf{0})}{2} \right), \quad t \in \mathbb{R}$$

Measuring roughness

GMM approach

Epilogue 000

Probabilistic setup

Log-normal spot volatility model

$$\sigma_t^2 := \xi \exp\left(\frac{Y_t - \frac{1}{2}\kappa(\mathbf{0})}{2}\right), \quad t \in \mathbb{R}$$

1. $(Y_t)_{t \in \mathbb{R}}$ is a stationary Gaussian process with mean zero and autocovariance function $\kappa(\cdot)$.



Probabilistic setup

Log-normal spot volatility model

$$\sigma_t^2 := \xi \exp\left(\frac{Y_t - \frac{1}{2}\kappa(0)}{2}\right), \quad t \in \mathbb{R}$$

- 1. $(Y_t)_{t \in \mathbb{R}}$ is a stationary Gaussian process with mean zero and autocovariance function $\kappa(\cdot)$.
- Collect the parameter ξ ∈ ℝ and all the parameters of κ(·) into θ ∈ Θ ⊂ ℝ^p.



Probabilistic setup

Log-normal spot volatility model

$$\sigma_t^2 := \xi \exp\left(\frac{\mathbf{Y}_t - \frac{1}{2}\kappa(\mathbf{0})}{2}\right), \quad t \in \mathbb{R}$$

- 1. $(Y_t)_{t \in \mathbb{R}}$ is a stationary Gaussian process with mean zero and autocovariance function $\kappa(\cdot)$.
- Collect the parameter ξ ∈ ℝ and all the parameters of κ(·) into θ ∈ Θ ⊂ ℝ^p.
- 3. Assume that Θ is compact with true θ_0 in its interior.



Covariance structure of integrated variance

Proposition

For any
$$\theta \in \Theta$$
, $k \in \mathbb{Z}$ and $\ell = 0, 1, ..., 1$. $g^{1}(\theta) := \mathbb{E}_{\theta}[IV_{k}] = \xi$

Covariance structure of integrated variance

Proposition

For any
$$\theta \in \Theta$$
, $k \in \mathbb{Z}$ and $\ell = 0, 1, ...,$
1. $g^{1}(\theta) := \mathbb{E}_{\theta}[IV_{k}] = \xi$
2. $g^{2}_{\ell}(\theta) := \mathbb{E}_{\theta}[IV_{k}IV_{k+\ell}] = \xi^{2} \int_{0}^{1} (1-y)(e^{\kappa(\ell+y)} + e^{\kappa(|\ell-y|)})dy$

Covariance structure of integrated variance

Proposition

For any
$$heta \in \Theta$$
, $k \in \mathbb{Z}$ and $\ell = 0, 1, \ldots$,

1.
$$g^1(heta) \coloneqq \mathbb{E}_{ heta}[IV_k] = \xi$$

2.
$$g_{\ell}^{2}(\theta) := \mathbb{E}_{\theta}[IV_{k}IV_{k+\ell}] = \xi^{2} \int_{0}^{1} (1-y)(e^{\kappa(\ell+y)} + e^{\kappa(|\ell-y|)})dy$$

If additionally $\lim_{u\to\infty} \kappa(u) = 0$ and some regularity conditions on κ hold, then:

3. $\operatorname{Cov}_{\theta}[IV_k, IV_{k+\ell}] \sim \operatorname{const}_{\xi,\kappa}\kappa(\ell)$ as $\ell \to \infty$

Measuring roughness

GMM approach

Epilogue 000

Stylised measurement error

Measurement setup

$$\widehat{IV}_k = IV_k + \varepsilon_k, \quad k \in \mathbb{Z}$$



Stylised measurement error

Measurement setup

$$\widehat{IV}_k = IV_k + \varepsilon_k, \quad k \in \mathbb{Z}$$

The stylised measurement errors ε_k , $k \in \mathbb{Z}$, should satisfy:

Stylised measurement error

Measurement setup

$$\widehat{IV}_k = IV_k + \varepsilon_k, \quad k \in \mathbb{Z}$$

The stylised measurement errors ε_k , $k \in \mathbb{Z}$, should satisfy:

1. $(IV_k, \varepsilon_k)_{k \in \mathbb{Z}}$ is a stationary and ergodic process under \mathbb{P}_{θ} for any $\theta \in \Theta$.
Measurement setup

$$\widehat{IV}_k = IV_k + \varepsilon_k, \quad k \in \mathbb{Z}$$

The stylised measurement errors ε_k , $k \in \mathbb{Z}$, should satisfy:

- 1. $(IV_k, \varepsilon_k)_{k \in \mathbb{Z}}$ is a stationary and ergodic process under \mathbb{P}_{θ} for any $\theta \in \Theta$.
- 2. $\theta \mapsto c(\theta) := \mathbb{E}_{\theta}[\varepsilon_1^2]$ is a finite-valued, continuous function on Θ .

Measurement setup

$$\widehat{IV}_k = IV_k + \varepsilon_k, \quad k \in \mathbb{Z}$$

The stylised measurement errors ε_k , $k \in \mathbb{Z}$, should satisfy:

- 1. $(IV_k, \varepsilon_k)_{k \in \mathbb{Z}}$ is a stationary and ergodic process under \mathbb{P}_{θ} for any $\theta \in \Theta$.
- 2. $\theta \mapsto c(\theta) := \mathbb{E}_{\theta}[\varepsilon_1^2]$ is a finite-valued, continuous function on Θ .
- 3. $\mathbb{E}_{\theta}[\varepsilon_{k}|\varepsilon_{k-1}, \varepsilon_{k-2}, \dots, Y] = 0$ for any $k \in \mathbb{Z}$ and any $\theta \in \Theta$.



Example (CLT approximation)

The central limit theorem (CLT) for realised variance (e.g., Fukasawa, 2010) says:

$$\sqrt{n}e_k^n = \sqrt{n}(RV_k^n - IV_k) \xrightarrow[n \to \infty]{d} \left(2\int_{k-1}^k \sigma_t^4 \mathrm{d}t\right)^{1/2} X_k$$

Example (CLT approximation)

The central limit theorem (CLT) for realised variance (e.g., Fukasawa, 2010) says:

$$\sqrt{n}e_k^n = \sqrt{n}(RV_k^n - IV_k) \xrightarrow[n \to \infty]{d} \left(2\int_{k-1}^k \sigma_t^4 \mathrm{d}t\right)^{1/2} X_k$$

where $X_k \sim N(0, 1)$, $k \in \mathbb{Z}$, are mutually independent and independent of Y.

Example (CLT approximation)

The central limit theorem (CLT) for realised variance (e.g., Fukasawa, 2010) says:

$$\sqrt{n}e_k^n = \sqrt{n}(RV_k^n - IV_k) \xrightarrow[n \to \infty]{d} \left(2\int_{k-1}^k \sigma_t^4 \mathrm{d}t\right)^{1/2} X_k$$

where $X_k \sim N(0, 1)$, $k \in \mathbb{Z}$, are mutually independent and independent of Y. Motivated by this result, we specify

$$\varepsilon_k := \left(\frac{2}{n} \int_{k-1}^k \sigma_t^4 \mathrm{d}t\right)^{1/2} X_k,$$

Example (CLT approximation)

The central limit theorem (CLT) for realised variance (e.g., Fukasawa, 2010) says:

$$\sqrt{n}e_k^n = \sqrt{n}(RV_k^n - IV_k) \xrightarrow[n \to \infty]{d} \left(2\int_{k-1}^k \sigma_t^4 \mathrm{d}t\right)^{1/2} X_k$$

where $X_k \sim N(0, 1)$, $k \in \mathbb{Z}$, are mutually independent and independent of Y. Motivated by this result, we specify

$$\varepsilon_k := \left(\frac{2}{n}\int_{k-1}^k \sigma_t^4 \mathrm{d}t\right)^{1/2} X_k, \quad \text{for which} \quad c(\theta) = \frac{2\xi^2}{n}e^{\kappa(0)}.$$



Moments

Covariance structure under measurement error

1.
$$\mathbb{E}_{\theta}[\widehat{IV}_{k}] = \mathbb{E}_{\theta}[IV_{k}] = g^{1}(\theta)$$



Moments

Covariance structure under measurement error

1.
$$\mathbb{E}_{\theta} \left[\widehat{IV}_{k} \right] = \mathbb{E}_{\theta} \left[IV_{k} \right] = g^{1}(\theta)$$

2.
$$\mathbb{E}_{\theta} \left[\widehat{IV}_{k} \widehat{IV}_{k+\ell} \right] = \begin{cases} \mathbb{E}_{\theta} \left[IV_{k}^{2} \right] + c(\theta) = g_{0}^{2}(\theta) + c(\theta), & \ell = 0\\ \mathbb{E}_{\theta} \left[IV_{k} IV_{k+\ell} \right] = g_{\ell}^{2}(\theta), & \ell > 0 \end{cases}$$



Moments

Covariance structure under measurement error

1.
$$\mathbb{E}_{\theta} \left[\widehat{IV}_{k} \right] = \mathbb{E}_{\theta} \left[IV_{k} \right] = g^{1}(\theta)$$

2.
$$\mathbb{E}_{\theta} \left[\widehat{IV}_{k} \widehat{IV}_{k+\ell} \right] = \begin{cases} \mathbb{E}_{\theta} \left[IV_{k}^{2} \right] + c(\theta) = g_{0}^{2}(\theta) + c(\theta), & \ell = 0\\ \mathbb{E}_{\theta} \left[IV_{k} IV_{k+\ell} \right] = g_{\ell}^{2}(\theta), & \ell > 0 \end{cases}$$

Moment selection

For a finite subset $\mathcal{L} \subset \{1,2,\ldots\}$ of lags,

$$egin{aligned} \widehat{\mathbb{IV}}_k &:= ig(\widehat{IV}_k, \widehat{IV}_k^2, \widehat{IV}_k \widehat{IV}_{k-\ell} : \ell \in \mathcal{L}ig), \quad k \in \mathbb{Z} \ \mathcal{G}_c(heta) &:= ig(g^1(heta), g_0^2(heta) + c(heta), g_\ell^2(heta) : \ell \in \mathcal{L}ig), \quad heta \in \Theta \end{aligned}$$



Generalised method of moments (GMM)

Estimating function

Define

$$\widehat{m}_N(heta) := rac{1}{N} \sum_{k=1}^N \widehat{\mathbb{IV}}_k - \mathcal{G}_c(heta), \quad heta \in \Theta,$$



Generalised method of moments (GMM)

Estimating function

Define

$$\widehat{m}_N(heta) := rac{1}{N} \sum_{k=1}^N \widehat{\mathbb{IV}}_k - \mathcal{G}_c(heta), \quad heta \in \Theta,$$

and note $\mathbb{E}_{\theta_0}[\widehat{m}_N(\theta_0)] = 0.$



Generalised method of moments (GMM)

Estimating function

Define

$$\widehat{m}_N(heta) := rac{1}{N} \sum_{k=1}^N \widehat{\mathbb{IV}}_k - \mathcal{G}_c(heta), \quad heta \in \Theta,$$

and note $\mathbb{E}_{\theta_0}[\widehat{m}_N(\theta_0)] = 0.$

GMM estimator

For a data-dependent weight matrix $W \in \mathbb{R}^{(2+|\mathcal{L}|) \times (2+|\mathcal{L}|)}$,

$$\widehat{ heta}_{N} := rgmin_{ heta \in \Theta} \widehat{m}_{N}(heta)' W \widehat{m}_{N}(heta)$$

The estimator is consistent and asymptotically normal under suitable conditions.



Specifying Y

(Rough) fractional stochastic volatility model (FSV/RFSV)

Let Y be a fractional Ornstein–Uhlenbeck process with Hurst index $H \in (0, 1)$,

$$Y_t =
u \int_{-\infty}^t e^{-\lambda(t-u)} \mathrm{d} W^H_u, \quad t \in \mathbb{R},$$



Specifying Y

(Rough) fractional stochastic volatility model (FSV/RFSV)

Let Y be a fractional Ornstein–Uhlenbeck process with Hurst index $H \in (0, 1)$,

$$Y_t =
u \int_{-\infty}^t e^{-\lambda(t-u)} \mathrm{d} W^H_u, \quad t \in \mathbb{R},$$

where:

• W^H is an fBm with the Hurst index H



Specifying Y

(Rough) fractional stochastic volatility model (FSV/RFSV)

Let Y be a fractional Ornstein–Uhlenbeck process with Hurst index $H \in (0, 1)$,

$$Y_t =
u \int_{-\infty}^t e^{-\lambda(t-u)} \mathrm{d} W^H_u, \quad t \in \mathbb{R},$$

where:

- W^H is an fBm with the Hurst index H
- $\nu > 0$ and $\lambda > 0$ are parameters



Specifying Y

(Rough) fractional stochastic volatility model (FSV/RFSV)

Let Y be a fractional Ornstein–Uhlenbeck process with Hurst index $H \in (0, 1)$,

$$Y_t =
u \int_{-\infty}^t e^{-\lambda(t-u)} \mathrm{d} W^H_u, \quad t \in \mathbb{R},$$

where:

- W^H is an fBm with the Hurst index H
- $\nu > 0$ and $\lambda > 0$ are parameters

Introduced as the fractional stochastic volatility (FSV) model by Comte and Renault (1998) with $H \in (\frac{1}{2}, 1)$.



Specifying Y

(Rough) fractional stochastic volatility model (FSV/RFSV)

Let Y be a fractional Ornstein–Uhlenbeck process with Hurst index $H \in (0, 1)$,

$$Y_t =
u \int_{-\infty}^t e^{-\lambda(t-u)} \mathrm{d} W^H_u, \quad t \in \mathbb{R},$$

where:

- W^H is an fBm with the Hurst index H
- $\nu > 0$ and $\lambda > 0$ are parameters

Introduced as the fractional stochastic volatility (FSV) model by Comte and Renault (1998) with $H \in (\frac{1}{2}, 1)$. Repurposed in the rough case $H \in (0, \frac{1}{2})$ by Gatheral, Jaisson and Rosenbaum (2018).



Application to simulated FSV/RFSV data

Setup

• Test both $\widehat{IV}_k := IV_k$ and $\widehat{IV}_k := RV_k$ (5 min).



- Test both $\widehat{IV}_k := IV_k$ and $\widehat{IV}_k := RV_k$ (5 min).
- *IV_k* without correction (*c*(θ) := 0), *RV_k* with CLT approximation correction and without.



- Test both $\widehat{IV}_k := IV_k$ and $\widehat{IV}_k := RV_k$ (5 min).
- IV_k without correction $(c(\theta) := 0)$, RV_k with CLT approximation correction and without.
- Use Gatheral, Jaisson and Rosenbaum's (2018) procedure to derive initial parameter guesses ("baseline").



- Test both $\widehat{IV}_k := IV_k$ and $\widehat{IV}_k := RV_k$ (5 min).
- *IV_k* without correction (*c*(θ) := 0), *RV_k* with CLT approximation correction and without.
- Use Gatheral, Jaisson and Rosenbaum's (2018) procedure to derive initial parameter guesses ("baseline").
- Following Fukasawa, Takabatake and Westphal (2022), consider *H* ∈ {0.05, 0.1, 0.3, 0.5, 0.7}.
- Other parameters set to keep the data realistic.
- 10000 replications with N = 4000.
- $\mathcal{L} := \{1, 2, 3, 5, 20, 50\}$, weight matrix W derived from Newey and West (1987) long-run covariance estimates.





















Setup

• Apply GMM estimation to 5-min daily realised variance data on all indices in the database (except Karachi SE 100 Index and Straits Times Index, due to gaps in time series).



- Apply GMM estimation to 5-min daily realised variance data on all indices in the database (except Karachi SE 100 Index and Straits Times Index, due to gaps in time series).
- Use CLT approximation correction.



- Apply GMM estimation to 5-min daily realised variance data on all indices in the database (except Karachi SE 100 Index and Straits Times Index, due to gaps in time series).
- Use CLT approximation correction.
- Other estimation settings as in the simulation experiment.



- Apply GMM estimation to 5-min daily realised variance data on all indices in the database (except Karachi SE 100 Index and Straits Times Index, due to gaps in time series).
- Use CLT approximation correction.
- Other estimation settings as in the simulation experiment.
- Inference (confidence intervals) using asymptotic normality.

N	1	ea	su	rii	ng	ro	ug	hn	ess	
С	1		0)				



Application to Oxford-Man realised variance data





Measuring the roughness of volatility

Generalised method of moments approach

Epilogue

22 / 24



Summary

• Volatility is rough, as far as realised variance is concerned.



Summary

- Volatility is rough, as far as realised variance is concerned.
- But inferring that spot volatility is rough is harder.



Summary

- Volatility is rough, as far as realised variance is concerned.
- But inferring that spot volatility is rough is harder.
- Integration (smoothes) and measurement error (roughens) become counteracting sources of bias.



Summary

- Volatility is rough, as far as realised variance is concerned.
- But inferring that spot volatility is rough is harder.
- Integration (smoothes) and measurement error (roughens) become counteracting sources of bias.
- But within a log-normal spot volatility model using GMM, we can accommodate both of these effects.


Concluding remarks

Summary

- Volatility is rough, as far as realised variance is concerned.
- But inferring that spot volatility is rough is harder.
- Integration (smoothes) and measurement error (roughens) become counteracting sources of bias.
- But within a log-normal spot volatility model using GMM, we can accommodate both of these effects.
- GMM estimation and inference suggest that spot volatility is under log-normality best described by a rough process.



Concluding remarks

Summary

- Volatility is rough, as far as realised variance is concerned.
- But inferring that spot volatility is rough is harder.
- Integration (smoothes) and measurement error (roughens) become counteracting sources of bias.
- But within a log-normal spot volatility model using GMM, we can accommodate both of these effects.
- GMM estimation and inference suggest that spot volatility is under log-normality best described by a rough process.
- Typically $\hat{H} \approx 0.02$, consistent with implied volatility calibrations.



Bibliography

M. Bennedsen, A. Lunde, and M. S. Pakkanen (2022): Decoupling the shortand long-term behavior of stochastic volatility. *Journal of Financial Econometrics* **20**(5), 961–1006.



A. E. Bolko, K. Christensen, M. S. Pakkanen and B. Veliyev (2023): A GMM approach to estimate the roughness of stochastic volatility. *Journal of Econometrics* **235**(2), 745–778.



- F. Comte and E. Renault (1998): Long memory in continuous-time stochastic volatility models. *Mathematical Finance* **8**(4), 291–323.
- M. Fukasawa (2010): Realized volatility with stochastic sampling. *Stochastic Processes and their Applications* **120**(6), 829–852.
- M. Fukasawa, T. Takabatake and R. Westphal (2022): Consistent estimation for fractional stochastic volatility model under high-frequency asymptotics. *Mathematical Finance* **32**(4), 1086–1132.



J. Gatheral, T. Jaisson and M. Rosenbaum (2018): Volatility is rough. *Quantitative Finance* **18**(6), 933–949.



W. K. Newey and K. D. West (1987): A simple, positive semi-definite, heteroscedasticity and autocorrelation consistent covariance matrix. *Econometrica* **55**(3), 703–708.