

**On the Limit Distributions of the First Moment of  
the High Boundary Crossing for Random Processes  
and Sequences**

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# 1. Introduction

Our aim - to give a simple derivation of the limit distribution of the first moment of the high boundary crossing (for wide range of random processes and sequences).

For historical reasons we start exposition with discrete-time sequences and then consider continuous-time processes.

Initial for us is the following well-known “*Anscombe Theorem*”.

Let  $X_1, X_2, \dots$  – i.i.d. random variables with

$$\mathbf{E}X_1 = a > 0, \quad \mathbf{E}|X_1 - a|^2 = \sigma^2 < \infty.$$

Denote  $S_n = X_1 + \dots + X_n$ ,  $S_0 = 0$ ,

and introduce the first-passage time moment

$$\tau_B = \min\{n : S_n \geq B\}, \quad B > 0.$$

For simplicity, we assume  $a > 0$  is fixed and  $B \rightarrow \infty$ .

Then with high probability  $\tau_B \approx B/a$ .

A more accurate result gives

**Theorem 1.** (Anscombe, 1952)

$$\mathcal{L} \left\{ \frac{\tau_B - B/a}{(\sigma^2 B/a^3)^{1/2}} \right\} \Rightarrow \mathcal{N}(0, 1) \quad B \rightarrow \infty. \quad (1)$$

(Was proved by (Heyde, 1966) and (Gut, 1974)).

We give simple proof of a more general variant of formula (1).

Let  $X_1, X_2, \dots$  – i.i.d.r.v. with  $\mathbf{E}X_1 = a$ ,  $0 < a < \infty$ .

**Assumption  $\mathbf{A}_1$ .** There exists a non-random sequence  $\{b_n\}$  and parameter  $\alpha \in (1, 2]$ , such that

$$\mathcal{L} \left\{ \frac{S_n - an}{b_n} \right\} \Rightarrow G_\alpha \quad n \rightarrow \infty, \quad (2)$$

where  $G_\alpha$  - stable distribution function with parameter  $\alpha$ .

It is known:

- 1) the limit distribution in (2) may be only stable distribution  $G_\alpha$ ;
- 2) if  $\alpha > 1$ , then  $b_n = Cn^{1/\alpha}$  with some constant  $C$ ;
- 3) distribution  $G_\alpha$  has any absolute moment of power  $b \in [0, \alpha)$ .

Denote by  $\xi_F$  random variable with distribution  $F$ .

**Theorem 2.** *For  $\tau_B$  the stochastic representation holds*

$$\tau_B = \frac{1}{a}(B - b_{B/a}\xi_{G_\alpha})[1 + o_B(1)]. \quad (3)$$

where  $o_B(1) \xrightarrow{\mathbf{P}} 0$ ,  $B \rightarrow \infty$ . In other words, for any fixed  $x$

$$\mathbf{P} \left\{ \frac{B - a\tau_B}{b_{B/a}} \leq x \right\} \rightarrow G_\alpha(x), \quad B \rightarrow \infty. \quad (4)$$

(Th. 2 was proved by other methods in (Heyde, 1967), (Gut, 1974)).

**Corollary 1.** *If  $\mathbf{E}|X_1 - a|^2 = \sigma^2$ ,  $0 < \sigma < \infty$ , then  $\alpha = 2$ ,  $b_n = \sigma n^{1/2}$ ,  $G_2(x) = \Phi(x)$  and “Anscombe Theorem” holds.*

Informally speaking, if we have a certain information on the behavior of  $\tau_B$ , then with a simple trick that information can be improved by one more level. For example, if we know that  $\tau_B \xrightarrow{\mathbf{P}} B/a, B \rightarrow \infty$  (see formula (6)), then it is possible to get limiting distribution of  $\tau_B$  (e.g. Theorem 1). Similarly, using the distribution of  $\tau_B$  (for example, formula (6)), it is possible to get next term (on  $B$ ) of limiting distribution of  $\tau_B$ .

## 2. Proof of Theorem 2

From assumption  $\mathbf{A}_1$  we have

$$S_n = an + b_n \xi_F + b_n o_n(1), \quad n \rightarrow \infty,$$

where  $o_n(1) \xrightarrow{\mathbf{P}} 0$  as  $n \rightarrow \infty$ .

By strong law of large numbers we can get (Heyde, 1966)

$$\frac{\tau_B}{B} \xrightarrow{a.s.} \frac{1}{a}, \quad B \rightarrow \infty, \quad (5)$$

and, in particular,

$$\tau_B = \frac{B[1 + o_B(1)]}{a}, \quad (6)$$

where  $o_B(1) \xrightarrow{\mathbf{P}} 0$  as  $B \rightarrow \infty$ .

Choose a fixed moment  $T = T(B)$  such that:

- 1) with high probability  $\tau_B > T$ ;
- 2)  $T$  is rather close to  $B/a$  (more exactly - below).

We show that finding the distribution of  $\tau_B$  is approx. equivalent to finding the distribution of  $B - S_T$  (for appropriate choice of  $T$ ).

Using strong Markovian property of  $\{S_n\}$ , moment  $\tau_B$  can be represented as

$$\begin{aligned}\tau_B &= (T + \tau_{B-S_T}) I_{\{\tau_B \geq T\}} + \tau_B I_{\{\tau_B < T\}} = \\ &= T + \tau_{B-S_T} - [(T + \tau_{B-S_T} - \tau_B) I_{\{\tau_B < T\}}].\end{aligned}$$

Assume that we can choose  $T$  such that  $[ ] \ll T + \tau_{B-S_T}$ . Then

$$\tau_B \approx T + \tau_{B-S_T}.$$

We set  $T = B(1 - \varepsilon)/a$  with small (but fixed)  $\varepsilon > 0$ . Then after standard algebra we get Theorem 2.  $\square$



The same approach can be applied to non-identically distributed r.v. or weakly dependent, etc.

### 3. Random Processes

The same approach can be applied to continuous-time case.

We limit ourselves to illustrative example:

$$\xi(t) = t + \sigma W(t), \quad t \geq 0,$$

$W(t)$ ,  $t \geq 0$ , – standard Wiener process;  $\sigma > 0$  – given constant.

Introduce the first-passage time moment

$$\tau_A = \min\{t > 0 : \xi(t) = A\}, \quad A > 0.$$

Characteristic function of  $\tau_A$ : for any  $\lambda$  random process

$$\eta(t) \stackrel{\text{def}}{=} \exp\left\{-\frac{\lambda^2\sigma^2 t}{2} + \lambda\sigma W(t)\right\} = \exp\left\{-\frac{\lambda^2\sigma^2 t}{2} - \lambda t + \lambda\xi(t)\right\}$$

is martingale. Therefore  $\mathbf{E}\eta(\tau_A) = 1$ , and then

(*Wald fundamental identity*)

$$\mathbf{E} \exp\left\{-\left(\frac{\lambda^2\sigma^2}{2} + \lambda\right)\tau_A\right\} = \exp\{-\lambda A\},$$

or, equivalently,

$$\mathbf{E} \exp\{\mu\tau_A\} = \exp\left\{\frac{2\mu A}{1 + \sqrt{1 - 2\mu\sigma^2}}\right\}.$$

After changing  $\mu = s/(\sigma\sqrt{A})$

$$\mathbf{E} \exp \left\{ s \frac{\tau_A - A}{\sigma\sqrt{A}} \right\} = \exp \left\{ \frac{2s^2}{\left[ 1 + \sqrt{1 - 2s\sigma/\sqrt{A}} \right]^2} \right\}$$

and then

$$\lim_{A \rightarrow \infty} \mathbf{E} \exp \left\{ s \frac{\tau_A - A}{\sigma\sqrt{A}} \right\} = \exp \left\{ \frac{s^2}{2} \right\}$$

Therefore we get

**Proposition 1.**

$$\mathcal{L} \left\{ \frac{\tau_A - A}{\sigma\sqrt{A}} \right\} \Rightarrow \mathcal{N}(0, 1), \quad A \rightarrow \infty. \quad (7)$$

We get the same result by simpler method. Clearly,

$$\frac{\tau_A}{A} \xrightarrow{a.s.} 1, \quad A \rightarrow \infty, \quad (8)$$

and, in particular,

$$\tau_A = A[1 + o_A(1)], \quad (9)$$

where  $o_A(1) \xrightarrow{\mathbf{P}} 0$  as  $A \rightarrow \infty$ .

Choose a fixed moment  $T = T(A)$  such that:

- 1) with high probability  $\tau_A > T$ ;
- 2)  $T$  is rather close to  $A$  (more exactly - below).

We show that finding the distribution of  $\tau_A$  is approx. equivalent to finding the distribution of  $A - \xi(T)$  (for appropriate choice of  $T$ ).

Again, moment  $\tau_A$  can be represented as

$$\begin{aligned}\tau_A &= [T + \tau_{A-\xi(T)}] I_{\{\tau_A \geq T\}} + \tau_A I_{\{\tau_A < T\}} = \\ &= T + \tau_{A-\xi(T)} - [(T + \tau_{A-\xi(T)} - \tau_A) I_{\{\tau_A < T\}}].\end{aligned}$$

Assume that we can choose  $T$  such that  $[ ] \ll T + \tau_{A-\xi(T)}$ . Then

$$\tau_A \approx T + \tau_{A-\xi(T)}.$$

We set  $T = A(1 - \varepsilon)$  with small (but fixed)  $\varepsilon > 0$ . Then after standard algebra we get Proposition 1.  $\square$