

# Statistical Properties of Affine Stochastic Delay Differential Equations

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# Outline

- 1 Stochastic Ordinary Differential Equations (SODEs)
- 2 Stochastic Delay Differential Equations (SDDEs)
- 3 Affine Stochastic Delay Differential Equations (ASDDEs)
- 4 Affine SDDEs with Fractional Derivatives

# Stochastic ordinary differential equations (SODEs)

$$dX(t) = \mu_\theta(X(t)) dt + \sigma_\theta(X(t)) dB(t), \quad t \in [0, T],$$
$$X(0) = \xi, \quad \text{independent of } B(\cdot), \theta \in \Theta \subseteq \mathbb{R}^k, \Theta \text{ open.}$$

The solution is a Markov process

Statistical problem:

- Estimate  $\theta$ , based on continuous time observation  $X^T = (X(t), t \in [0, T])$ .
- Study asymptotic properties for  $T \rightarrow \infty$ .

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- Estimate  $\theta$ , based on continuous time observation  $X^T = (X(t), t \in [0, T])$ .
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Assumptions:

- $\sigma_{\theta}(x) = \sigma(x)$ , independent of  $\theta$ ,
- $(X(t), t \geq 0)$  ergodic,
- $\xi$  has the invariant distribution.

Kutoyants (2004), [12]

# Maximum likelihood method

$$\begin{aligned} \ell_T &:= \log \frac{d\mathbb{P}_\theta^T}{d\mathbb{P}_{\theta_0}^T} \\ &= \int_0^T \frac{\mu_\theta(X(s)) - \mu_{\theta_0}(X(s))}{\sigma^2(X(s))} dX(s) - \frac{1}{2} \int_0^T \frac{(\mu_\theta(X(s)) - \mu_{\theta_0}(X(s)))^2}{\sigma^2(X(s))} ds, \end{aligned}$$

$$\hat{\theta}_T = \arg \max_{\theta \in \Theta} \ell_T(\theta, (X(t), t \in [0, T])).$$

$\mathbb{P}_\theta^T$  ... distribution of  $(X(t), t \leq T)$  on  $C([0, T])$ .

# Method of localization

$$\eta = \theta + \phi_T(\theta)u,$$

where  $u \in \mathbb{R}^k$ ,  $\phi_T(\theta)$   $k \times k$ -matrix, Ibragimov, Khasminskii (1981), [9].

$$\phi_T(\theta) \longrightarrow 0, \quad \text{for } T \rightarrow \infty. \quad (\text{normalizing matrix})$$



## Method of localization

$$\eta = \theta + \phi_T(\theta)u,$$

where  $u \in \mathbb{R}^k$ ,  $\phi_T(\theta)$   $k \times k$ -matrix.

Under some conditions of regularity:

The family  $(\mathbb{P}_\theta^T, \theta \in \Theta)$  is **locally asymptotically normal (LAN)** at every  $\theta \in \Theta$ . It means

$$Z_{\theta,T}(u) := \log \frac{d\mathbb{P}_{\theta+\phi_T(\theta)u}^T}{d\mathbb{P}_\theta^T} \xrightarrow{T \rightarrow \infty} \langle u, \Delta(\theta) \rangle - \frac{1}{2} \langle I(\theta)u, u \rangle,$$

where

$$\Delta(\theta) \sim N(0, I(\theta)) \quad \text{and} \quad \phi_T(\theta) = T^{-1/2} J_k,$$

$$I(\theta) := \mathbb{E}_\theta \left( \frac{\dot{\mu}_\theta(\xi) \dot{\mu}_\theta^*(\xi)}{\sigma(\xi)} \right).$$

## Method of localization

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where  $u \in \mathbb{R}^k$ ,  $\phi_T(\theta)$   $k \times k$ -matrix.

It is well known, that in the non regular case other limits of the log-likelihood functions are possible, e.g., [LAMN](#), [LAQ](#) cases and even

$$Z_\theta(u) = \exp \left\{ C_\theta B^H(u) - \frac{C_\theta^2}{2} |u|^{2H} \right\},$$

where  $B^H(\cdot)$  is the [fractional Brownian motion with Hurst parameter](#)  $H \in [1/2, 1]$ ,  $C_\theta$  is constant.



## The affine Markovian case (Ornstein-Uhlenbeck case)

$$dX(t) = \theta X(t) dt + dB(t), \quad t \in [0, T],$$

$$X(0) = \xi, \quad \text{independent of } B(\cdot).$$

$$\implies X(t) = \xi e^{\theta t} + \int_0^t e^{\theta(t-s)} dB(s), \quad t \geq 0,$$

$$\ell_T = \log L_T = \log \frac{d\mathbb{P}_\theta^T}{d\mathbb{P}_0^T} = \theta \int_0^T X(t) dX(t) - \frac{\theta^2}{2} \int_0^T X^2(t) dt,$$

$$\hat{\theta} = \underbrace{\left( \int_0^T X^2(t) dt \right)^{-1}}_{I_T} \underbrace{\int_0^T X(t) dX(t)}_{V_T} = I_T^{-1} V_T.$$



# The affine Markovian case (Ornstein-Uhlenbeck case)

$$\hat{\theta}(T) - \theta = \left( \int_0^T X^2(t) dt \right)^{-1} \int_0^T X(t) dB(t) = I_T^{-1} V_T^0.$$

Asymptotic properties of  $\hat{\theta}(T)$  for  $T \rightarrow \infty$ .

Three different cases:

$\theta < 0$ : LAN,

$\theta > 0$ : LAMN,

$\theta = 0$ : LAQ.

The Markov property is paid by the assumption that the coefficient  $\mu(X(t))$  and  $\sigma(X(t))$  depend on the present state  $X(t)$  of  $X$  only, i.e. that  $X$  has no memory.

In many cases there is some time between reason and effect of a phenomenon

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## Biology and Population dynamics:

- time to maturity
- time to educate
- incubation period

## Economics: time between

- taking a decision and occurrence of the feedback
- time to build, to transport, to store

The are different attempts to include memory effects into modeling real phenomena:

- CARMA process (Brockwell, 2001, [2])
- continuous time GARCH processes (Brockwell, Chadraa, Lindner, 2006, [1])

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# Stochastic delay differential equations (SDDEs)

$$dX(t) = \mu_\theta(X_t) dt + \sigma_\theta(X_t) dB(t), \quad t \geq 0, \quad (1)$$

$$X_0 = (\xi(s), s \in [-r, 0]), \quad (1')$$

- $X_0$  - **initial process**, independent of  $B$ ,
- $X_t = (X(t+s), s \in [-r, 0])$ : "segment of  $X(\cdot)$  at time  $t$ ",
- $r \geq 0$  length of the **memory**,
- $\mu_\theta(x), \sigma_\theta(x) : [0, \infty) \times C([-r, 0]) \rightarrow \mathbb{R}^1, \theta \in \Theta \subseteq \mathbb{R}^k$ .

The solutions of (1) and (1') are not Markov processes.

No explicit solutions, even in simple cases.

- Numerics of SDDEs was treated by Buckwar et al., [3], [4].
- The case of infinitely long memory ( $r = \infty$ ) has some particular aspects, see e.g. Riedle, [16].



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## Deterministic DDEs

$$\begin{aligned}\frac{dx(t)}{dt} &= f(t, x_t), \quad t \geq 0, \\ x_0 &= \xi \in C([-r, 0]).\end{aligned}$$

- Widely developed theory: see e.g. Hale, Lunel, [8].

## Stochastic DDEs

- Mao (1997, 2007), [13],
- Mohammed (1984), [14].



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# Affine stochastic delay differential equations (ASDDEs)

$$dX(t) = \int_{-r}^0 X(t+s)a(ds)dt + dB(t), \quad t \geq 0, \quad (2)$$

$$X(u) = \xi(u), \quad u \in [-r, 0]. \quad (2')$$

$a(\cdot)$  finite signed measure.

Autoregressive schema with continuous time.

## Example (Ornstein-Uhlenbeck case)

$$a(ds) = a\mathbf{1}_{\{0\}}(ds),$$

$$dX(t) = aX(t)dt + dB(t),$$

$$X(0) = \xi.$$

# The solution of the affine equation

$$dX(t) = \int_{-r}^0 X(t+s)a(ds)dt + dB(t), \quad t \geq 0, \quad (2)$$

$$X(u) = \xi(u), \quad u \in [-r, 0], \quad (2')$$

admits the representation

$$X(t) = \xi(0)x_0(t) + \int_{-r}^0 \int_u^0 \xi(s)x_0(t+u-s)ds a(du) + \int_0^t x_0(t-s)dB(s).$$



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$x_0(\cdot)$  denotes the so called **fundamental solution**,

$$\dot{x}_0(t) = \int_{-r}^0 x_0(t+s)a(ds),$$

$$x_0(s) = \mathbf{1}_{\{0\}}(u), \quad u \in [-r, 0].$$





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### Example ( Ornstein-Uhlenbeck case)

$$x_0(t) = \exp[at],$$

$$X(t) = \xi \exp[at] + \int_0^t \exp[a(t-s)]dB(s).$$

# The fundamental solution $x_0(t)$ :

$$\dot{x}_0(t) = \int_{-r}^0 x_0(t+s)a(ds),$$
$$x_0(t) = \mathbf{1}_{\{0\}}(t), \quad t \in [-r, 0].$$

The fundamental solution  $x_0(\cdot)$

- has no explicit expression,
- may be oscillating instead of being monotone.



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$$x_0(t) = \mathbf{1}_{\{0\}}(t), \quad t \in [-r, 0].$$

## Laplace-transform

$$\hat{x}_0(\lambda) = \int_0^{\infty} e^{-\lambda s} x_0(s) ds, \quad \operatorname{Re} \lambda > \|a\|,$$

$$\hat{x}_0(\lambda) = \left( \lambda - \int_{-r}^0 e^{\lambda s} a(ds) \right)^{-1}, \quad \operatorname{Re} \lambda > \|a\|.$$

# The fundamental solution $x_0(t)$ :

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$$x_0(t) = \mathbf{1}_{\{0\}}(t), \quad t \in [-r, 0].$$

## Characteristic function

The denominator

$$h(\lambda) = \lambda - \int_{-r}^0 e^{\lambda s} a(ds), \quad \lambda \in \mathbb{C},$$

is called the **characteristic function** of the deterministic differential equation.

$h(\cdot)$  is holomorphic on  $\mathbb{C}$ .

# The fundamental solution $x_0(t)$ :

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## Characteristic function

The denominator

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is called the **characteristic function** of the deterministic differential equation.

$$v_0 := \max\{\operatorname{Re}\lambda \mid h(\lambda) = 0\} < \infty \quad (\text{top coefficient}),$$

$$v_{i+1} := \max\{\operatorname{Re}\lambda \mid h(\lambda) = 0, \operatorname{Re}\lambda < v_i\}, \quad i \geq 0.$$

## The fundamental solution $x_0(t)$ :

$$\dot{x}_0(t) = \int_{-r}^0 x_0(t+s)a(ds),$$

$$x_0(t) = \mathbf{1}_{\{0\}}(t), \quad t \in [-r, 0].$$

The inverse Laplace transform yields for some  $c < v_0$ :

$$x_0(t) = p(t) \exp\{v_0 t\} + o(e^{ct}), \quad t \geq 0.$$

$p(t)$  polynomially bounded function of  $t$ .

$v_0 < 0$ :  $x_0$  tends **exponentially** to zero.



# Back to the stochastic case

$$dX(t) = \int_{-r}^0 X(t+s)a(ds)dt + dB(t), \quad t \geq 0 \quad (2)$$

$$X(u) = \xi(u), \quad t \in [-u, 0]. \quad (2')$$

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- There exists a stationary solution if and only if  $v_0 < 0$ .
- This stationary solution is centered, Gaussian with covariance function

$$K(h) = \int_0^\infty x_0(s+h)x_0(s)ds, \quad h \geq 0.$$

- It satisfies the analogue of the **Yule-Walker equation**

$$\dot{K}(t) = \int_{-r}^0 K(t+u)a(du), \quad t > 0.$$



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- There exists a stationary solution if and only if  $v_0 < 0$ .
- This stationary solution is centered, Gaussian with covariance function

$$K(-h) = K(h) = \int_0^{\infty} x_0(s+h)x_0(s)ds, \quad h \geq 0.$$

- If  $v_0 < 0$ , then the covariance function decreases **exponentially** to zero as  $\exp[v_0 h]$  for  $h \rightarrow \infty$ .

”Short-range dependence”

# Statistical questions:

$$dX(t) = \int_{-r}^0 X(t+s)a(ds) dt + dB(t), \quad t \geq 0,$$

$$X(u) = \xi(u), \quad u \in [-r, 0].$$

Likelihood process:

$$L_t = \frac{d\mathbb{P}_a^t}{d\mathbb{P}_0^t},$$

Log-likelihood process:

$$\ell_t = \log L_t = \underbrace{\int_{-r}^0 \int_0^t X(s+u) dX(s) a(du)}_{V_t} - \frac{1}{2} \underbrace{\int_{-r}^0 \int_{-r}^0 \int_0^t X(s+u) X(s+v) ds a(du) a(dv)}_{I_t}$$

$$= \langle V_t, a \rangle - \frac{1}{2} \langle I_t a, a \rangle \rightarrow \max.$$

Obviously

- $V_t(\cdot) \in C([-r, 0])$  and  $I_t(\cdot, \cdot) \in C([-r, 0]^2)$ ,
- $I_t$  defines a Hilbert-Schmidt operator from  $BV([-r, 0])$  to  $C([-r, 0])$  and we call it the **Fisher information operator**.

## The maximum likelihood equation

$$I_t \hat{a}_t(\cdot) = V_t(\cdot), \quad \text{i.e.,}$$

$$\int_{-r}^0 \left( \int_0^t X(s+\cdot) X(s+u) ds \right) \hat{a}_t(du) = \int_0^t X(s+\cdot) dX(s).$$

$\hat{a}_t$  ... nonparametric maximum likelihood estimator.

$I_t$  is a compact linear operator, thus the ML problem is an ill-posed problem.

**Nonparametric case:** Nonparametric estimation for  $a(\cdot)$  based on the continuous observation of  $(X(t), t \in [0, T])$  see Reiss (2002), [15].

**Parametric case:**

$$dX(t) = \int_{-r}^0 X(t+s) a_{\theta}(ds) dt + dB(t), \quad t \geq 0, \quad (3)$$

$$X(s) = \xi(s), \quad s \in [-r, 0]. \quad (3')$$

## Aim

- Estimate  $\theta$  based on continuous observation  $X(t), t \in [0, T]$ .
- Study asymptotic properties of the estimator.

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## Aim

- Estimate  $\theta$  based on continuous observation  $X(t), t \in [0, T]$ .
- Study asymptotic properties of the estimator.

## Example (Gushchin, K. (1999), [5])

$$dX(t) = [aX(t) + bX(t-r)] dt + dB(t),$$

$$\theta = (a, b) \in \mathbb{R}^2.$$

Maximum likelihood estimator  $\hat{\theta}(T)$ :

$$I_T^{-1} \cdot V_T = \begin{pmatrix} \int_0^T X^2(t) dt & \int_0^T X(t)X(t-r) dt \\ \bullet & \int_0^T X^2(t-r) dt \end{pmatrix}^{-1} \begin{pmatrix} \int_0^T X(t) dX(t) \\ \int_0^T X(t-r) dX(t) \end{pmatrix}$$

and

$$\hat{\theta}_T - \theta = I_T^{-1} V_T^0 \quad \text{with}$$

$$V_T^0 = \begin{pmatrix} \int_0^T X(s) dB(s) \\ \int_0^T X(s-r) dB(s) \end{pmatrix}.$$

## Example (Gushchin, K. (1999), [5])

$$dX(t) = [aX(t) + bX(t-r)] dt + dB(t),$$

$$\theta = (a, b) \in \mathbb{R}^2.$$

There are **eleven different cases** of asymptotic behaviour of  $\hat{\theta}_T$ , depending on  $\theta = (a, b)^T$ :

- in one case N: LAN-property,
- in three cases  $M_1, M_2, M_3$ : LANM-property,
- in five cases  $Q_1, Q_2, \dots, Q_5$ : LAQ-property,
- in two cases  $P_1, P_2$ : PLAMN-property.

Both, the corresponding sets of parameters and the limit distributions are calculated explicitly.

In this basic example: LAN-property holds iff  $v_0 < 0$ , i.e., iff there exists a stationary solution.

## The LAN-property

It arises the question, under which general assumptions on  $\theta \mapsto a_\theta$  the LAN-property holds.

$$dX(t) = \int_{-r}^0 X(t+s) a_\theta(ds) dt + dB(t), \quad (3)$$

$$X(u) = \xi(u), \quad u \in [-r, 0]. \quad (3')$$

- $\theta \in \Theta \subseteq \mathbb{R}^k$ ,  $\Theta$  open,
- $M_S = \{a : (3) \text{ admits a stationary solution}\}$ ,
- Assume  $\mathcal{A} := \{a_\theta(ds) : \theta \in \Theta \subseteq \mathbb{R}^k\} \subseteq M_S$ .



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Localization around a fixed  $\theta \in \Theta$ :

$$\eta = \theta + \varphi_T(\theta)u, \quad u \in \mathbb{R}^k$$

where  $\varphi_T(\theta)$  is a  $k \times k$ -regular matrix, with

$$\varphi_T(\theta) \rightarrow 0 \text{ for } T \rightarrow \infty \quad (\text{normalizing matrix})$$

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### Assumption (A): (Condition on regularity)

- 1)  $\exists C < \infty : \|a_\theta - a_\eta\|_V \leq C \|\theta - \eta\|$ ,
- 2)  $\theta \mapsto \int_{[-r,0]} g(u) a_\theta(du)$ , continuously differentiable in  $\Theta$  with the gradient  $\int_{[-r,0]} g(u) \dot{a}_\theta(du)$  for all  $g \in C[-r,0]$ ,  $\dot{a}_\theta = (\dot{a}_\theta^{(1)}, \dot{a}_\theta^{(2)}, \dots, \dot{a}_\theta^{(k)})^T$ ,  $\theta \in \Theta$ ,
- 3)  $\mathcal{A} = (a_\theta, \theta \in \Theta)$ ,  $\Theta$  open, bounded subset of  $\mathbb{R}^k$
- 4)  $a_\theta \in M_S$ ,  $\theta \in \bar{\Theta}$ ,
- 5)  $a_\theta \neq a_\eta$ ,  $\theta \in \Theta$ ,  $\eta \in \bar{\Theta}$ ,  $\theta \neq \eta$ .

## Example

$$dX(t) = \theta X(t-r)dt + dB(t).$$

Assumption (A) holds

$$\int_{-r}^0 g(u) a_{\theta}(du) = \theta g(-r),$$

$$a_{\theta}(\cdot) = \theta \delta_{-r}, \quad \dot{a}_{\theta}(\cdot) = \delta_{-r}.$$

## Example

$$dX(t) = bX(t-\theta)dt + dB(t)$$

does not satisfy assumption (A) because

$$\theta \mapsto \int_{[-r,0]} g(u) a_{\theta}(du) = bg(-\theta)$$

is not differentiable for all  $g \in C[-r,0]$ .

## Theorem (Gushchin, K. (2003), [6])

Under the assumption (A) for every compact set  $K \subseteq \Theta$  it holds uniformly in  $\theta \in K$ :

- 

$$\sqrt{T}(\hat{\theta}_T - \theta) \xrightarrow{D} N(0, \Sigma^{-1}(\theta)), \quad T \longrightarrow \infty,$$

where

$$\Sigma_{ij}(\theta) = \int \int_{\mathcal{I} \times \mathcal{I}} K_{\theta}(u-v) \dot{a}_{\theta,i}(du) \dot{a}_{\theta,j}(dv).$$

- All moments of  $\sqrt{T}(\hat{\theta}_T - \theta)$  under  $\mathbb{IP}_T^{\theta}$  tend as  $T \longrightarrow \infty$  to the corresponding moments of a normal distribution with parameters  $(0, \Sigma^{-1}(\theta))$ .
- The maximum likelihood estimator  $\hat{\theta}_T$  is asymptotically efficient in  $K$ .

As an intermediate step it was proved

$$\begin{aligned}
 Z_{T,\theta}(u) &= \frac{d\mathbb{P}_{\theta+\varphi_T u}^T}{d\mathbb{P}_{\theta}^T} \\
 &= \exp\left\{ \int_0^T \int_{-r}^0 X(s+v) (a_{\theta+\varphi_T u}(dv) - a_{\theta}(dv)) dX(s) \right. \\
 &\quad \left. - \frac{1}{2} \int_0^T \left( \int_{-r}^0 \int_{-r}^0 X(s+v) X(s+w) \left[ a_{\theta+\varphi_T u}(dv) a_{\theta+\varphi_T u}(dw) \right. \right. \right. \\
 &\quad \left. \left. \left. - a_{\theta}(dv) a_{\theta}(dw) \right] \right) ds \right\} \\
 &\xrightarrow{T \rightarrow \infty} \exp\left\{ \langle Z, u \rangle - \frac{\|u\|^2}{2} \right\}, \quad (\text{Local asymptotic normality})
 \end{aligned}$$

where  $Z \sim N(0, J_k)$ .

## Example

$$dX(t) = bX(t - \theta)dt + dB(t), \quad b < 0, \text{ fixed,}$$

$$\theta \in \left(0, \frac{1}{e|b|}\right) =: \Theta \quad (\text{ensures } \nu_0 < 0).$$

$$\begin{aligned} Z_{T,\theta}(u) &:= \frac{d\mathbb{P}_{\theta+\varphi_T u}}{d\mathbb{P}_\theta} \\ &= \exp\left(b \int_0^T (X(t-\theta-\varphi_T u) - X(t-\theta)) dX(t) \right. \\ &\quad \left. - \frac{b^2}{2} \int_0^T (X(t-\theta-\varphi_T u) - X(-\theta))^2 dt\right). \end{aligned}$$



## Lemma (K., Kutoyants (2000), [11])

For  $\varphi_T = b^2 T^{-1}$  the marginal distributions of  $Z_{T, \theta + \varphi_T u}(\cdot)$  converge to the marginal distributions of

$$Z(u) = \exp \left\{ \tilde{B}(u) - \frac{1}{2}|u|^2 \right\}, \quad u \in \mathbb{R}^1,$$

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Remember:

$$a_\theta(dv) = b \delta_{-\theta}(dv):$$

$$Z_T(u) = \frac{d\mathbb{P}_{\theta + \varphi_T(\theta)u}^T}{d\mathbb{P}_\theta^T} \longrightarrow Z(u) = \exp[\tilde{B}(u) - \frac{1}{2}|u|].$$

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$$Z_T(u) = \frac{d\mathbb{P}_{\theta + \varphi_T(\theta)u}^T}{d\mathbb{P}_\theta^T} \longrightarrow Z(u) = \exp[uB - \frac{1}{2}u^2].$$





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uniformly over every compact set  $K \subseteq \mathbb{R}^1$ .

## Lemma (Gushchin, K. (2008), [7])

For every  $H \in [\frac{1}{2}, 1]$  one can find examples  $a_\theta(dv)$  with

$$\frac{d\mathbb{P}_{\theta + \varphi_T(\theta)u}^T}{d\mathbb{P}_\theta^T} \longrightarrow Z(u) = \exp \left[ B^H(u) - \frac{\mathbb{E}(B^H(u))^2}{2} \right],$$

where  $B^H$  is a fractional Brownian motion with Hurst parameter  $H$ .

$H = \frac{1}{2}$  :  $B^H(u)$  Brownian motion

$H = 1$  :  $B^H(u) = uB$ ,  $B \sim N(0, 1)$ .

- 1 Stochastic Ordinary Differential Equations (SODEs)
- 2 Stochastic Delay Differential Equations (SDDEs)
- 3 Affine Stochastic Delay Differential Equations (ASDDEs)
- 4 Affine SDDEs with Fractional Derivatives**

# Fractional calculus: an introduction

## Riemann-Liouville Integral

$$J^\alpha f(x) = \frac{1}{\Gamma(\alpha)} \int_0^x (x-t)^{\alpha-1} f(t) dt.$$

## Riemann-Liouville Derivative

$$D^\alpha f := D^1 J^{1-\alpha} f, \quad \alpha \in (0, 1).$$

## Caputo Derivative

$$D_C^\alpha f := D^\alpha [f - f(0)].$$

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## Fractional affine DDE

$$D_0^\alpha y(t) = \int_{-r}^0 y(t+s)a(ds) + f(t), \quad (5)$$

$$y(s) = \xi(s), \quad s \in [-r, 0], \quad \xi(\cdot) \in C([-r, 0]). \quad (5')$$

Solution of (5, 5'):

$$y(t) = \xi(0)x_0(t) + D^{1-\alpha} \int_{-r}^0 \left( \int_u^0 x_0(t+u-v)\xi(v)dv \right) a(du) + \int_0^t x_0(t-s)f(s)ds.$$

Here  $x_0(\cdot)$  denotes the **fundamental solution** of (5, 5'), i.e.,

$$\begin{aligned} D_0^\alpha x_0(t) &= \int_{-r}^0 x_0(t+s)a(ds), \\ x_0(s) &= \mathbf{1}_{\{0\}}(s), \quad s \in [-r, 0]. \end{aligned}$$

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## Fundamental solution

$$D_0^\alpha x_0(t) = \int_{-r}^0 x_0(t+s)a(ds),$$

$$x_0(s) = \mathbf{1}_{\{0\}}(s), \quad s \in [-r, 0].$$

Laplace-transform:

$$\hat{x}_0(\lambda) = \frac{\lambda^{\alpha-1}}{\lambda^\alpha - \int_{-r}^0 e^{\lambda s} a(ds)}, \quad \operatorname{Re} \lambda > \|a\|^{\frac{1}{\alpha}}.$$

Characteristic function:

$$h_\alpha(\lambda) = \left[ \lambda^\alpha - \int_{-r}^0 e^{\lambda s} a(ds) \right] \cdot \lambda^{1-\alpha} = \lambda - \lambda^{1-\alpha} \int_{-r}^0 e^{\lambda s} a(ds).$$





## Fundamental solution

$$D_0^\alpha x_0(t) = \int_{-r}^0 x_0(t+s) a(ds),$$

$$x_0(s) = \mathbf{1}_{\{0\}}(s), \quad s \in [-r, 0].$$

There is a qualitative jump from  $\alpha = 1$  to  $\alpha < 1$  :  
The characteristic function

$$h_\alpha(\lambda) = \lambda - \lambda^{1-\alpha} \int_{-r}^0 e^{\lambda s} a(ds)$$

is

- holomorphic only on  $\mathbb{C} \setminus \mathbb{R}_-$ ,
- discontinuous on  $\mathbb{R}_-$ ,



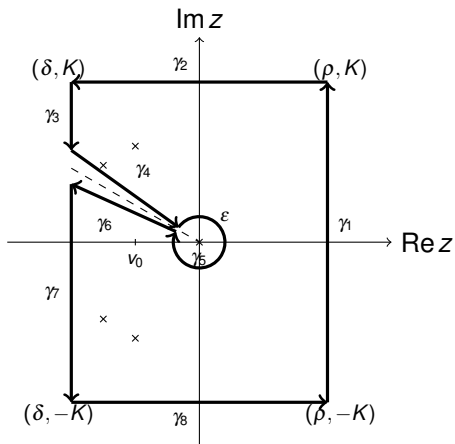


Figure: Contour integration of  $\chi_\lambda^{-1}(z)e^{zt}$ .

## Fundamental solution

$$D_0^\alpha x_0(t) = \int_{-r}^0 x_0(t+s)a(ds),$$

$$x_0(s) = \mathbf{1}_{\{0\}}(s), \quad s \in [-r, 0].$$

## Theorem (Krol (2008), [10])

For  $v_0 < 0$  we have  $x_0(t) = p(t)e^{v_0 t} + g(t)$ ,  $t \geq 0$ ,

with  $p(t)$  polynomially bounded function and

$$g(t) = \begin{cases} O(t^{-\alpha}), & \text{if } a([-r, 0]) \neq 0, \\ 1 + O(t^{-k_0 + \alpha}), & \text{if } a([-r, 0]) = 0, \text{ for some } k_0 \in \mathbf{N}. \end{cases}$$

Thus the asymptotic of the fundamental solution  $x_0(t)$  for  $t \rightarrow \infty$  is **not exponentially** but **polynomially**.



## Stochastic case

$$X(t) = \frac{1}{\Gamma(\alpha)} \int_0^t (t-s)^{\alpha-1} \int_{-r}^0 X(s+u) a(du) ds + B(t),$$

$$X(s) = \xi(s), \quad s \in [-r, 0].$$

## Theorem (Krol (2008), [10])

*There exists a stationary solution iff  $v_0 < 0$ ,  $a([-r, 0]) \neq 0$  and  $\alpha > \frac{1}{2}$ .*

## Stochastic case

$$X(t) = \frac{1}{\Gamma(\alpha)} \int_0^t (t-s)^{\alpha-1} \int_{-r}^0 X(s+u) a(du) ds + B(t),$$

$$X(s) = \xi(s), \quad s \in [-r, 0].$$

## Corollary

The covariance function  $K(t)$  of the stationary solution is given by







$$K(t) = \int_0^t x_0(s) x_0(s+t) ds.$$

$K(t)$  tends polynomially to zero for  $t \rightarrow \infty$ .

Thus,  $(X(t), t \geq 0)$  has

”long-range dependence”

Thank you for your attention!

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